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Cointegration Analysis of Exchange Rate Volatility and Agricultural Exports in Turkey: an Ardl Approch

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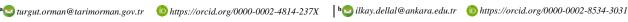
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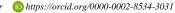
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Keywords: Volatility ARDL bounds testing Producer price index Agricultural exports Real effective exchange rate ABSTRACT

This study aims to reveal the impact of exchange rate volatility on agricultural exports of Turkey by using the Autoregressive Distributed Lag Model. While quarterly time series data covering period of 2001: Q1 to 2018: Q4 were used to carry out analyses, Exponential Generalized Autoregressive Conditional Heteroscedasticity (1.1) is used to acquire exchange rate volatility series. The research findings showed that agricultural export is cointegrated with exchange rate volatility, producer price index and real effective exchange rate. Furthermore, our findings indicate that increases in real effective exchange rate have a statistically significant positive influence on the export volume whereas exchange rate volatility has negative impact on it.









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Introduction

Agricultural exports have been playing a crucial role in national economies in terms of its contribution to the productivity of domestic production and income for the people involved in agriculture sector. Besides, it is a major source of foreign exchange that is necessary to finance imports and pave the way for development of other sectors particularly in developing countries (FAO, 1995). Turkey as a developing country, is not an exception to this fact. Agricultural export including food products has a share of nearly 10% in total exports of Turkey. Furthermore, Turkey is the leading exporter of many agricultural products such as nut, apricot, grapes, fig, quince and wheat flour in the world markets (Turkstat, 2020).

In recent years, agricultural exports and its relation with exchange rates have been an important area of policy discussions. Although the issue is on the stage for many transition economies throughout the world, Turkish case is more valuable to evaluate in terms of the effect of exchange rate and its volatility on the agricultural exports due to the unprecedented and intense movements in Turkish currency markets since the second quarter of the 2013.

There is a considerable interest in the literature to measure the effect of exchange rate and its volatility on the agricultural trade flows both for developed and developing countries. However, existing studies have not reached consensus on the significancy and direction of the effect. Some studies found that exchange rate volatility is an important determinant of Turkish agricultural export (Erdal et al., 2012, Yazici 2012), while others showed that exchange rate and its variability do not have significant impact on export (Buguk et al., 2003).

The results mainly vary acording to the selected period of time, variables used in the model and the countries studied. In this context, this study empirically investigates the impact of exchange rate volatility on the agricultural exports of Turkey employing Autoregressive Distributed

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Lag Model (ARDL) developed by Pesaran et al., 2001. While quarterly time series data covering period of 2001: Q1 to 2018: Q4 were used to carry out the analyses, Exponential Generalized Autoregressive Conditional Heteroscedasticity (EGARCH) (1.1) is employed to obtain exchange rate volatility series.

A Brief Overview of Agricultural Exports and Exchange Rate Markets in Turkey

There are three main classifications issued by the United Nations (UN) to record and measure trade data. Agricultural trade data is measured by selecting specific chapters and/or codes listed in those classifications (TURKSTAT, 2008). According to the International Standard Industrial Classification of All Economic Activities (ISIC) Rev. 3, agricultural exports are measured by summing chapters 01, 02 and 05-agriculture, hunting and forestry and fishing. Food industry products are not included in those chapters. Some selected chapters (01-24) from Harmonized System (HS) which is one of the detailed classifications used by World Trade Organization (WTO), Food and Agriculture Organization (FAO) and any national institutions to report agricultural export data, are generally used to define agricultural exports including food industry product (UNSD, 2008). Some selected chapters from Standard International Trade Classification (SITC) are also used to measure agricultural export data (UNSD,

Turkish agricultural exports were 6 billion USD in 2018 according to ISIC. On the other hand, it was 19.3 billion US Dollar (USD) when food industry products are included according to HS classification. Share of agricultural exports also varies according to classification selected. It is 3-4% according to ISIC and 10-11% according to HS/SITC (Table 1), (Turkstat,2020).

Main products exported in 2018 were nuts and fruits, citrus fruits or peel of melons (12.8%), preparations of vegetables, fruits or other parts of plants (6.3%) malt, starches, gluten, inulin and products of the millings industry (3.9%), (Turkstat, 2020).

Major export markets for Turkish agriculture in 2018 were Iraq (14.9%), Germany (7.1%), USA (7.7%), Russian Federation (4.6%) and Italy (3.6%). It can be observed that there is a quite high concentration in agricultural export in terms of products and trade partners. The share of first three products in total agricultural exports was 23% and share of first five trade partners was 35% in 2018 (Turkstat, 2020).

Turkey implemented a fixed exchange rate policy like many other developing countries until 1980. The policy has evolved from fixed exchange rate regime to flexibly managed exchange rate regime from that time onwards. Turkey has a floating exchange rate regime since 2001 and there are limited interventions by Central Bank of the Republic of Turkey (CBRT) to prevent excessive volatility in the currency markets (Aytac, 2016). Exchange rates, as it is one of the instrumental policy tools, have the ability to influence trade flows in open economies, especially in those which implement flexible regime. It affects the export, import, and prices of products worldwide.

As it can be seen from the Figure 1, there is an upward trend in Turkish currency markets since mid-2013. Turkish Lira is depreciating against US Dollar and Euro. This is

calculated as one of the highest depreciations among the developing countries (Figure 1).

Real Effective Exchange Rate, which is obtained from CBRT database, is a basic indicator to measure the international competition power of Turkey against its trading partners. It is weighted average of the prices in Turkey relative to the prices of its main trade partners in international markets. A downward trend has been observed between 2001-2018 (Figure 2), (CBRT, 2020).

Exchange rate volatility is an important factor that has to be taken into account especially for developing countries such as Turkey. It has generally been assumed that volatility, meaning uncertainty can impose cost on the trade and may negatively affect the exports for risk-averse traders. However, there is no consensus on the relation between exports and volatility in the literature.

Data and Methodology

The empirical analysis is carried out using quarterly observations covering the period of 2001: Q1 – 2018: Q4. Deseasonalized real agricultural export (AGEXS) data obtained from TurkStat database is used for agricultural exports. The classification ISIC Rev.3 is selected. CPI based real effective exchange rate (REER) data, average quarterly value of the daily conditional heteroscedasticity for USD/TL exchange rate volatility (Kim, 2017) (REERV) which is obtained through EGARCH (1.1) and producer price index of agricultural products (AGPPI) are used in the model as explanatory variables. Log form of all series except exchange rate volatility are used in order to avoid excessive changes and having linear values (Ozer, 2013).

Engle (1982) and Bollerslev (1986) have stated that generalized autoregressive conditional heteroskedastic (GARCH) models are instrumental if we model the timevarying volatility of the time series. The conventional GARCH model allows the conditional variance to be dependent upon previous own lags. Exponential Generalized Autoregressive Conditional Heteroscedasticity (EGARCH) (1.1) introduced by Nelson (1991) is one of the techniques to determine the volatility in REER and to calculate REERV series. EGARCH models along with a conditional variance formulation are evaluated that they capture asymmetric response in the conditional variance successfully and had been proved superior to other asymmetric conditional variance in literature. As the non-negativity constraint in GARCH model can be violated by the estimated method, EGARCH is found more useful since variance will be positive although parameters are negative (Su, 2010).

Findings of the EGARCH (1.1) model are shown in Table 2 and indicate that all parameters are statistically significant at 1% level of probability. Among all, α gives insight into the persistency of volatility and in case it is statistically significant, size of a shock has a significant impact on volatility. Meanwhile γ gives insight into how the sign of a shock has an influence on the future volatility. If γ is statistically significant and its coefficient is smaller than zero, it triggers a leverage effect where the volatility of a negative shock has greater effect than that of a positive shock (Table 2).

Autoregressive Distributed Lag Model (ARDL), which is developed by Pesaran and Shin (1999) and Pesaran et al. (2001) has been performed to analyse impact of exchange rate and its volatility on agricultural exports. It relies on estimation of an Unrestricted Error Correction Model

(UECM) which is widely accepted as superior to other cointegration techniques. It can estimate short and long run coefficients simultaneously and is useful for small sample size. Besides, it can be used for variables composed of I (0) and I (1) levels (Srinivasan and Kalaivani, 2012).

ARDL model specification used in the paper is;

ARDL-UECM

 $\Delta lnagexs_{t} = \alpha_{0} + \sum_{i=1}^{p} \beta_{1} \Delta lnreer_{t-i} + \sum_{i=1}^{p} \beta_{2} \Delta lnreer_{t-i} + \sum_{i=1}^{p} \beta_{3} \Delta lnagppi_{t-i} + \sum_{i=1}^{p} \beta_{4} \Delta lnagexs_{t-i} + \gamma_{1} lnreer_{t-1} + \gamma_{2} lnreer_{t-1} + \gamma_{2} lnagppi_{t-1} + \gamma_{4} lnagexs_{t-1} + \varepsilon_{t}$ (Equation 1)

• Long-Run Relation

 $lnagexs_{t} = a_{0} + \sum_{i=1}^{m} a_{1} \Delta lnreer_{t-i} + \sum_{i=1}^{n} a_{2} \Delta lnreerv_{t-i} + \sum_{i=1}^{n} a_{3} \Delta lnagppi_{t-i} + \sum_{i=1}^{q} a_{4} \Delta lnagexs_{t-i} + \eta_{t}$ (Equation 2)

• Short-Run Dynamics

 $\Delta lnagexs_t = b_0 + \sum_{i=1}^{m} b_1 \Delta lnreer_{t-i} + \sum_{i=1}^{n} b_2 \Delta lnreer_{t-i} + \sum_{i=1}^{y} b_3 \Delta lnagepi_{t-i} + \sum_{i=1}^{q} b_3 \Delta lnagexs_{t-i} + \prod ECM_{t-1} + \psi_t$ (Equation 3)

Table 1. Turkish Agricultural Exports (Billion US Dollar)

= 110 1				
	HS (0-24)	ISIC Rev3 (01, 02 and 05)	SITC (Selected Codes)	Total Exports
2014	18.0	6.4	18.7	157.6
2015	16.8	6.1	17.4	143.8
2016	16.2	5.8	16.9	142.5
2017	16.9	5.7	17.3	157.0
2018	19,2	6	18,1	168

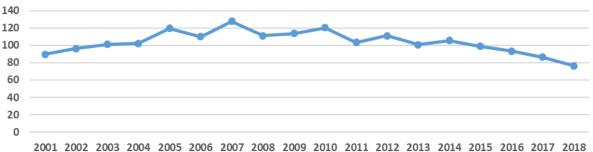
Source: TurkStat, Author's Calculations.

Table 2. EGARCH modelling of REER series

Water Francisco					
		Variance Equ	lation		
	Coefficient	Std. Error	z-Statistic	Probability	
ω	-2.816557	0.120350	-23.40308	0.0000	
α	0.261654	0.008810	29.69871	0.0000	
γ	0.062283	0.005165	12.05945	0.0000	
β	0.610182	0.017094	35.69482	0.0000	



Figure 1. USD and EURO against Turkish Lira Source: CBRT.



2003 2004 2005 2006 2007 2008 2009 2010 2011 2012 2013 2014 2015 2016 2017 2018 Figure 2. Real Effective Exchange Rate (CPI 2003=100)

Source: CBRT.

Results and Discussion

Augmented Dickey Fuller (ADF) unit root test results together with diagnostic and stability tests of ARDL are discussed in this section at first. Empirical findings derived from the aforementioned methodology to find the relation between real effective exchange rate, conditional heteroscedasticity for USD/TL exchange rate, producer price index and agricultural exports are described in the subsequent section.

Preliminary Analyses

We initially employ the standard ADF unit root test to find out at which level variables are stationary. Results are presented in the Table 3.

Diagnostic Tests and Stability of the ARDL Process

Breusch-Godfrey Serial Correlation LM test and Breusch-Pagan-Godfrey Heteroscedasticity test were implemented in order to prove reliability of the ARDL model. Additionally, CUSUM and CUSUM of squares tests were implemented to examine its stability. As a result, null hypothesizes of both LM and heteroscedasticity tests were accepted, meaning that errors do not exhibit autocorrelation and they are homoscedastic (Table 4). Results of CUSUM and CUSUM of squares tests, which are presented in Figure 3 shows that there is no instability in the equation during the sample period.

ARDL Cointegration Analyses

Findings of ARDL Bounds test for cointegration are presented in Table 5. Bounds test result shows that the F-statistic (5.94) exceeded the upper bound critical value at one percent significant level. Null hypothesis stating that there is no cointegration is rejected. It is confirmed that long-term equilibrium relationship is existed between Turkey's Agricultural export and USD/TL exchange rate volatility, agricultural producer price index and real effective exchange rate. Furthermore, cointegrated variables means that there is an adjustment process in the short run.

ADRL Long-run Estimates

As it is proved that independent and dependent variables are cointegrated, equation (2) which is selected according to Akaike Information Criterion (ARDL-

3,0,0,3) was estimated to obtain coefficients for long-run relationship (Table 6).

Results show that the impact of real effective exchange rate (lnREER) and producer price index (AGPPI) on agricultural exports of Turkey are positive and significant in the long run. Estimated coefficient of the exchange rate volatility (lnREERV) is negative but unable to secure statistical significance. More specifically, the results indicated that %1 increase in lnREER and lnAGPPI led to a 0,63% and 0,74% increase in export volume respectively.

ARDL Short-Run Estimates

Equation (3) was estimated in order to obtain short-run dynamic coefficients of long-run relationship and its findings are presented in Table 6. The error correction model coefficient, which indicate speed of adjustment from short run to long run, is statistically significant and negative. This means that when Turkish agricultural export is deviating from long-run equilibrium level, 43% of that deviation is eliminated at each quarter of the year and around 7 months is needed for a full adjustment towards the long run equilibrium level.

Findings in the Table 7 show that the impact of reel effective exchange rate on agricultural exports of Turkey is positive and significant in the short run, meaning that higher reel effective exchange rate tends to increase agricultural exports in short run. Coefficients of exchange rate volatility and producer price index were negative but unable to secure statistical significance (Table 7).

Positive correlation between real effective exchange rate and export is prominent finding of our estimation whereas most studies in the literature suggest negative correlation (Kafle and Kennedy (2012)), (Cheung and Sengupta (2013)). On the other hand, an IMF Working Paper developed by Iossifov and Fei (2019) proved that export of Turkey increased during the period of appreciation of Turkish Lira. Much of this result is attributed to increasing total factor productivity of Turkey compared with its trading partners especially between 2000-2008. Inelastic demand of agricultural goods and Gross domestic product (GDP) of importing countries might be other factors that cause increasing export while REER is going up.

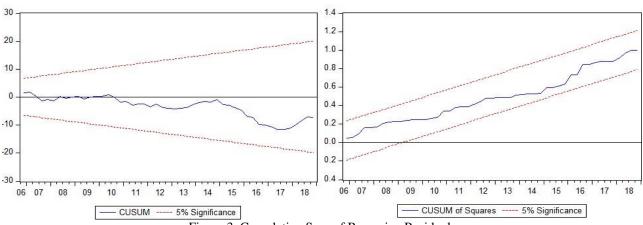


Figure 3. Cumulative Sum of Recursive Residuals

Table 3. Results of Unit Root Test

		Level		1st difference	
	Intercept	Intercept-Trend	Intercept	Intercept-Trend	
AGEXS	-1.889595	-0.352469	-3.642520*	-	I(1)
REER	-1.415140	-1.618434	-8.612701*		I(1)
REERV	-8.660283*				I(0)
AGPPI	3.517416**	-4.278305*			I(0)

Note: *,***,*** represents rejection of null hypothesis at 1, 5, 10 % significance level. The variables are composed of I(0) and I(1).

Table 4. Diagnostic Tests

	Breusch-Godfrey S	Serial Correlation LM Test		
F-statistic	2.014563	Prob. F(2.54)	0.1433*	
Obs*R-squared	4.790864	Prob. Chi-Square(2)	0.0911*	
	Heteroscedasticity T	est: Breusch-Pagan-Godfrey		
F-statistic 0.937262 Prob. F(12.56) 0.5176				
Obs*R-squared 11.54031 Prob. Chi-Square(12) 0.4833*				
Scaled explained SS	9.420060	Prob. Chi-Square(12)	0.6667	

Table 5. Bounds F-Test

	Computed F-Statistic: 5.94* k=4	
Significance	Lower Bound	Upper Bound
10%	2.45	3.52
5%	2.86	4.01
2.5%	3.25	4.49
1%	3.74	5.06

Notes: * indicates statistic exceeds upper bonds value at 1 percent significance level.

Table 6. Long-Run Coefficients of the ARDL (3,0,0,3)

Dependent Variable: lnAGEXS				
Regressor	Coefficient	t-statistic	Prob.Value	
Constant	3.725244	4.114341	0.0001	
lnREER	0.638663*	2.657931	0.0102	
InREERV	-0.001833	-0.753474	0.4543	
lnAGPPI	0.741602*	8.439292	0.0000	

Notes: * (**) represent significance at one and five percent level respectively.

Table 7. Error Correction Coefficients of the ARDL (3,0,0,3)

Dependent Variable: /	∆lnAGEXS	
Coefficient	t-statistic	Prob.Value
3.725244	4.114341	0.0001
0.253943	2.062005	0.0439
0.100883	0.819242	0.4161
0.209919	1.918714	0.0601
0.277981**	2.332797	0.0233
-0.000798	-0.725530	0.4711
-0.060596	-0.189089	0.8507
-0.187319	-0.414788	0.6799
0.110550	0.245234	0.8072
0.460151	1.544850	0.1280
-0.435256	-5.005021	0.0000
	Coefficient 3.725244 0.253943 0.100883 0.209919 0.277981** -0.000798 -0.060596 -0.187319 0.110550 0.460151	3.725244 4.114341 0.253943 2.062005 0.100883 0.819242 0.209919 1.918714 0.277981** 2.332797 -0.000798 -0.725530 -0.060596 -0.189089 -0.187319 -0.414788 0.110550 0.245234 0.460151 1.544850

Notes: * (**) represents significance at one and five per cent level respectively.

In contrast to relation between export and REER, our estimation is resulted in negative correlation between export and exchange rate volatility both in short and long run. While results of studies carried out by Thuy (2019) and Kafle and Kennedy (2015) are in agreement with our findings, Tunalioğlu et al. (2013) and Erdal et al. (2012) found positive relationship between agricultural export and exchange rate volatility. Different results might be

obtained depending on time period of analyses, sector and economic development level of examined country. In the light of our findings, it could be interpreted that having a sustainable and stable exchange rate may help to enhance real agricultural export. Additionally, it might be wise to improve capacity of Turkish exporting companies in terms of using hedging instruments which will help to cope with adverse currency movements.

Conclusion

This study aims to reveal the impact of real effective exchange rate, USD/TL exchange rate volatility, producer price index of agricultural products on real agricultural exports of Turkey. For this purpose, ARDL methodology is implemented with quarterly data covering the period 2001: Q1 – 2018: Q4. Besides, volatility data has been obtained by using EGARCH (1.1) model.

The selected ARDL (3,0,0,3) model, which is proved valid by diagnostic and stability tests, present that there is a long-run cointegration relationship among the variables.

The study results confirm that impact of real effective exchange rate (lnREER) on agricultural

exports of Turkey are positive and significant in both short and long run, implying that higher reel effective exchange rate tends to increase real exports. This may partly be explained by higher total factor productivity of Turkish agricultural sector compared with world average and importing countries.

On the other hand, estimated coefficient of the exchange rate volatility (lnREERV) was found to be negative and insignificant in both short and long run. This could be interpreted as adverse currency movements might be threat for Turkish export revenues in agriculture sector. Therefore, putting in place measures that will prevent unexpected currency movements may help to avoid risks associated with currency. In order to draw conclusions that generalize impact of exchange rate and its volatility on export, each country and each product need to be analyzed and it is necessary to take into account factors such as GDP, total factor productivity and demand elasticity of traded products. Such considerations are important next steps in the future analysis.

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